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Correlates of Cognitive Change

Timothy A. Salthouse University of Virginia

Although there has been considerable interest in identifying potential correlates of cognitive change, results of past studies have been inconsistent. The present study incorporated a number of methodological features intended to maximize sensitivity to detect characteristics of individuals with different amounts of cognitive change. Cognitive change in 5 cognitive abilities was analyzed with 2nd-order latent growth curve models applied to data from a moderately large sample of healthy adults ranging from 18 to 99 years of age (*N*s of 4,802 with 1 occasion, 2,265 with 2 occasions, and 1,128 with 3 occasions). There was significant individual difference variance in the longitudinal changes in several cognitive abilities, even in separate analyses of individuals between 18 years of age and 39, between 40 and 64, and 65 and over. Potential correlates of change included measures of self-rated health, vision, mood, personality, and lifestyle. Most of the potential correlates of change had high reliability, and several analyses were based on even more reliable factors determined by the variance common to multiple measures. Despite favorable conditions for detecting correlates of change, there was little evidence that cognitive change was moderated by any of the variables examined. Possible reasons for the inconsistent results regarding correlates of cognitive change are discussed.

Keywords: longitudinal, aging, moderation

Because variables found to have significant correlations with cognitive change may be informative about the factors contributing to successful and unsuccessful aging, and perhaps even provide clues about the mechanisms involved in longitudinal change, there has been a great deal of interest in identifying correlates of the average level, and of the magnitude of change, in cognitive functioning in healthy adults. In fact, because of the potential to enhance quality of life in old age and possibly prolong the period of independent living, [Hendrie et al. \(2006,](#page-19-0) p. 13) suggested that "identification of factors that can help people maintain or enhance cognitive or emotional health becomes a major public health goal."

A relatively large number of variables have been found to be correlated with measures of cognitive functioning in crosssectional comparisons, but results from cross-sectional studies only indirectly reflect change, and they do not allow analyses of individual differences in change, because the age comparisons are derived from different people. The focus in the present report is on research investigating predictors of change in longitudinal studies of cognitive functioning. Of primary interest was the identification of characteristics of people with different patterns of cognitive change. Because there have been a number of recent reviews of the literature (e.g., [Bielak, 2010;](#page-19-1) [Clouston et al., 2013;](#page-19-2) [Daffner, 2010;](#page-19-3) [Daviglus et al., 2010;](#page-19-4) [Depp, Vahia, & Jeste, 2010;](#page-19-5) [Hendrie et al.,](#page-19-0)

[2006;](#page-19-0) [Hertzog, Kramer, Wilson, & Lindenberger, 2008;](#page-19-6) [Miller,](#page-20-0) [Taler, Davidson, & Messier, 2012;](#page-20-0) [Plassman, Williams, Burke,](#page-20-1) [Holsinger, & Benjamin, 2010;](#page-20-1) [Sofi et al., 2011\)](#page-21-0), only a limited number of articles not included in the earlier reviews are discussed below. In order to organize the coverage, potential correlates of change are grouped into seven broad categories consisting of demographic characteristics, health, sensory ability, mood, personality and disposition, self-efficacy, and lifestyle.

Potential Correlates of Change

Demographic characteristics such as age, sex, and education have frequently been found to be related to cognition, and therefore they are important variables to control when analyzing relations of other variables with cognitive change. More negative cognitive change at older ages has been reported in many studies (e.g., [Lamar, Resnick, & Zonderman, 2003;](#page-20-2) [Mitchell et al., 2012;](#page-20-3) [Parisi et al., 2011;](#page-20-4) [Rönnlund & Nilsson, 2006;](#page-20-5) [Van Dijk, Van](#page-21-1) [Gerven, Van Voxtel, Van der Elst, & Jolles, 2008\)](#page-21-1). A few studies have reported differential change in males and females (e.g., [Parisi](#page-20-4) [et al., 2011\)](#page-20-4), but there are also numerous reports of no sex differences in change (e.g., [Finkel, Reynolds, McArdle, Gatz, &](#page-19-7) [Pedersen, 2003;](#page-19-7) [Lamar et al., 2003;](#page-20-2) [Lövdén et al., 2004\)](#page-20-6) or mixed patterns in different cognitive variables (e.g., [Mitchell et al., 2012;](#page-20-3) [Van Dijk et al., 2008\)](#page-21-1). With respect to education, [Hendrie et al.](#page-19-0) [\(2006,](#page-19-0) p. 22) concluded that "higher levels of education were almost uniformly reported to be protective for both cognitive and emotional outcomes." Although it is true that some studies have found less decline among individuals with higher levels of education (e.g., [Parisi et al., 2011\)](#page-20-4), other studies have found a relation of education with change only in some cognitive variables (e.g., [Singh-Manoux et al., 2011\)](#page-21-2) or have not found a relation between education and cognitive decline (e.g., [Glymour, Tzourio, & Du](#page-19-8)[fouil, 2012;](#page-19-8) [Karlamangla et al., 2009;](#page-20-7) [Mitchell et al., 2012;](#page-20-3)

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Correspondence concerning this article should be addressed to Timothy A. Salthouse, Department of Psychology, University of Virginia, 102 Gilmer Hall, Charlottesville, VA 22904. E-mail: salthouse@virginia.edu

[Tucker-Drob, Johnson, & Jones, 2009;](#page-21-3) [Van Dijk et al., 2008;](#page-21-1) [Zahodne et al., 2011\)](#page-21-4).

Health is a plausible correlate of cognitive change because a number of health conditions are known to affect the level of cognitive functioning, and at least some of them could be associated with more rapid cognitive decline. Health status has been assessed in a variety of different ways, including various types of physical examinations, as well as counts of medications, diseases, or illness episodes. Because they are easy to obtain, the most common measures of health are subjective ratings of one's health status. Although extremely simple, self-ratings of health have been found to be correlated with mortality (e.g., [Idler & Benyamini,](#page-20-8) [1997;](#page-20-8) [Miilunpalo, Vuori, Oja, Pasanen, & Urponen, 1997;](#page-20-9) [Singh-](#page-21-5)[Manoux et al., 2007\)](#page-21-5), physician visits [\(Miilunpalo et al., 1997\)](#page-20-9), and various biomarkers [\(Jylhä, Volpato, & Guralnik, 2006\)](#page-20-10).

Some studies have found poorer self-rated health to be associated with greater cognitive decline (e.g., [Carmelli, Swan, LaRue,](#page-19-9) [& Eslinger, 1997;](#page-19-9) [Gold et al., 1995;](#page-19-10) [Van Hooren et al., 2005;](#page-21-6) [Wahlin, Maitland, Backman, & Dixon, 2003\)](#page-21-7), but other studies have found different patterns for different variables (e.g., [Meijer,](#page-20-11) [van Boxtel, van Gerven, van Hooren, & Jolles, 2009;](#page-20-11) [Van Dijk et](#page-21-1) [al., 2008\)](#page-21-1) or have not found relations of self-rated health with cognitive change (e.g., [Anstey, Hofer, & Luszcz, 2003;](#page-19-11) [Hultsch,](#page-20-12) [Hertzog, Small, & Dixon, 1999;](#page-20-12) [Small, Dixon, & McArdle, 2011\)](#page-21-8).

Because relevant information cannot be processed if it is not adequately registered, sensory ability could also be a factor moderating cognitive change. Indeed, several studies have reported significant correlations between change in sensory function and change in cognitive functioning (e.g., [Anstey et al., 2003;](#page-19-11) [Linden](#page-20-13)[berger & Ghisletta, 2009;](#page-20-13) [Newson & Kemps, 2005;](#page-20-14) [Sternäng,](#page-21-9) [Jonsson, Wahlin, Nyberg, & Nilsson, 2010\)](#page-21-9).

A relatively large number of studies have examined relations of mood on cognitive change. Three major hypotheses have been proposed regarding the relation between mood and cognitive functioning. One is that negative mood is not a cause of cognitive decline but instead is a consequence of awareness of cognitive declines. A second hypothesis is that negative mood and cognitive change are both attributable to a third factor, such as vascular disease or reduced frontal lobe activity. The third hypothesis is most relevant to the issue of moderators of cognitive change because it postulates that negative mood influences subsequent cognitive change, perhaps because negative mood is associated with high levels of cortisol, which contribute to dysregulation of the hypothalamic– pituitary–adrenal axis, with negative consequences for hippocampal integrity and memory. There is consensus in the reviews that more depressive symptomatology is associated with more rapid cognitive decline (e.g., [Daviglus et al., 2010;](#page-19-4) [Hendrie et al., 2006;](#page-19-0) [Hertzog et al., 2008\)](#page-19-6). Significant relations of level of depressive symptoms to change in cognitive functioning have been reported in several recent studies (e.g., [Bielak, Gerstorf, Kiely, Anstey, &](#page-19-12) [Luszcz, 2011;](#page-19-12) [Köhler et al., 2010;](#page-20-15) [Van Den Kommer et al., 2013\)](#page-21-10), but other studies have not found relations between baseline depressive symptoms and subsequent cognitive change (e.g., [Jajodia](#page-20-16) [& Borders, 2011;](#page-20-16) [Mortensen, Barefoot, & Avlund, 2012\)](#page-20-17).

There is a moderately large literature documenting relations between aspects of personality and level of cognitive functioning at a single point in time (e.g., [Soubelet & Salthouse, 2011;](#page-21-11) [von](#page-21-12) [Stumm & Ackerman, 2013\)](#page-21-12). Personality and disposition have also been examined as potential correlates of change, in part because

these characteristics could affect the amount and type of activity one pursues (e.g., [Soubelet & Salthouse, 2010;](#page-21-13) [von Stumm &](#page-21-12) [Ackerman, 2013\)](#page-21-12). Results with these variables have not been very consistent, as significant relations of neuroticism on cognitive change have been reported in some studies (e.g., [Chapman et al.,](#page-19-13) [2012\)](#page-19-13) but not in others (e.g., [Sharp, Reynolds, Pedersen, & Gatz,](#page-21-14) [2010\)](#page-21-14), and no effect of openness or other personality traits on change in cognitive ability was found in a recent study by [Hogan,](#page-20-18) [Staff, Bunting, Deary, and Whalley \(2012\).](#page-20-18)

Because less negative change might be expected among individuals with a more positive outlook regarding their own level of cognition, a few studies have examined relations of self-efficacy to cognitive change. As noted by [Hertzog et al. \(2008\),](#page-19-6) the findings in this area have been mixed, although it should be noted that two recent studies reported significant correlations between change in subjective assessments of memory and change in objectively assessed memory (i.e., [Mascherek & Zimprich, 2011;](#page-20-19) [Parisi et al.,](#page-20-4) [2011\)](#page-20-4).

A large number of studies have examined relations between aspects of lifestyle and cognitive change. Many different types of lifestyle activities have been examined, but only cognitive activities and physical activities are considered here. There has been enormous variation in how cognitive activities have been evaluated, as the assessments have ranged from presence or absence of participation in one or several activities, to the total number of activities in which one was engaged in a specified period, and to the number of hours per week engaged in activities classified as cognitively stimulating. Methodological issues associated with assessment of cognitive activity have been discussed by a number of reviewers (e.g., [Bielak, 2010;](#page-19-1) [Ghisletta, Bickel, & Lövdén,](#page-19-14) [2006;](#page-19-14) [Hultsch et al., 1999;](#page-20-12) [Salthouse, 2006,](#page-20-20) [2010;](#page-20-21) [Salthouse,](#page-20-22) [Berish, & Miles, 2002\)](#page-20-22), including the almost complete lack of information about the validity of the activity reports.

Reviewers of the literature on cognitive activity and cognitive change have differed in their interpretations of the results. For example, [Hertzog et al. \(2008,](#page-19-6) p. 22) suggested that, "overall, these data strongly support the hypothesis that a higher level of engagement in mentally stimulating activity is associated with reduced loss of cognition in old age." In contrast, [Daviglus et al. \(2010\)](#page-19-4) were more cautious in stating that "limited but inconsistent evidence suggests that increased involvement in cognitive activities in later life may be associated with slower cognitive decline and lower risk for mild cognitive impairment." Recent studies have also been mixed, as some significant correlations between change in activity and change in cognition were reported in [Small, Dixon, McArdle, and Grimm \(2012\),](#page-21-15) but no relations of activity with cognitive change were reported in two other studies (e.g., [Bielak, Anstey, Christensen, & Windsor, 2012;](#page-19-15) [Mitchell et al., 2012\)](#page-20-3).

As with the assessment of cognitive activity, there has been considerable variation in how physical activity has been assessed. For example, the evaluations have ranged from the presence or absence of any activity, to frequency of engagement in specific activities such as gardening or sailing, to estimates of metabolic expenditures across specific activities in metabolic equivalent of task units based on frequency, duration, and intensity (see [Miller et](#page-20-0) [al., 2012\)](#page-20-0). Several reviewers have noted the weak validity of self-reports of physical activity (e.g., [Atienza et al., 2011;](#page-19-16) [Prince et](#page-20-23) [al., 2008;](#page-20-23) [Shephard, 2003\)](#page-21-16), which may be attributable to influences of social desirability [\(Adams et al., 2005\)](#page-19-17), memory limitations in the remembering frequency and duration of activities, and acrossperson variability in the interpretation of the nature of physical activity. Another parallel with the research on cognitive activity is discrepant interpretations of the existing evidence by reviewers. For example, [Miller et al. \(2012\)](#page-20-0) claimed that "the association between exercise and preserved cognition during aging is clearly demonstrated," and [Sofi et al. \(2011\)](#page-21-0) stated that the results "suggest a significant and consistent protection for all levels of physical activity against the occurrence of cognitive decline." In contrast, other reviewers qualified their conclusions by suggesting that the evidence was *growing* [\(Hendrie et al., 2006,](#page-19-0) p. 24) or was *preliminary* [\(Daviglus et al., 2010,](#page-19-4) p. 180). Results of recent studies have also been mixed, as [Clouston et al. \(2013\)](#page-19-2) found a correlation of physical activity at baseline with longitudinal change in cognition, but [Lindwall et al. \(2012\)](#page-20-24) reported a relation of baseline physical activity with change only in a verbal fluency measure and not in other cognitive measures.

Methodological Considerations

This brief overview indicates that each category of potential correlate of cognitive change has had inconsistent results. Furthermore, two reviews incorporating formal guidelines to evaluate the nature of the evidence concluded that the overall quality of evidence was low [\(Daviglus et al., 2010;](#page-19-4) [Plassman et al., 2010\)](#page-20-1). Future research investigating correlates of cognitive change should therefore incorporate as many desirable methodological features as possible. For example, the measurement of potential correlates should be sensitive, reliable, and valid. In addition, because the number of possible correlates is very large, instead of considering them separately and treating them as if they were all independent, relations among the variables should be identified to determine if the correlates form meaningful dimensions of individual differences, and if so, analyses should be carried out on measures of these dimensions and not simply on individual variables. Unique influences should also be investigated by considering groups of potential correlates simultaneously, rather than separately and independently.

The cognitive assessment should include several cognitive domains, with multiple indicators of each domain to emphasize effects on cognitive abilities rather than on individual variables that include test-specific influences and measurement error. As with the assessment of potential correlates, the measurement of cognitive functioning should be sensitive and reliable, with no restrictions attributable to measurement floors or ceilings. Furthermore, evidence of measurement invariance across occasions is desirable to ensure that any change is primarily quantitative and not qualitative. Measures of general cognition could be examined, but they may obscure differences in relations across cognitive domains, and there is little advantage of the enhanced reliability often associated with aggregate variables if the measures in each ability domain have high reliabilities.

The analytical methods should be sensitive to effects on change distinct from effects on level, which is not necessarily the case with all methods, such as those based on difference scores. In addition, age, sex, and years of education should be included in the analyses to control for influences of these variables when examining the relations of primary interest. It is also important to

consider variability of the measures of change in both the correlates and the cognitive variables because the critical factor affecting relations with other variables is not the magnitude of change but instead the amount of reliable variance in change. That is, if there is little evidence of differential change, in the form of significant individual difference variance in the measures of change, one cannot expect to identify correlates of differences in change that do not exist [\(Hertzog & Nesselroade, 2003\)](#page-19-18). An indirect indication of variability in change can be obtained from stability coefficients because high stability implies little variability in change. However, low stability is not sufficient to infer variability in change, because a low stability coefficient could be a consequence of low reliability, and thus both short-term reliability and stability need to be considered in evaluating variance of change. Change variance can also be estimated directly with statistical models, such as the latent growth model employed in the present study.

Statistical power to detect possible differences in cognitive change also needs to be considered in studies investigating correlates of cognitive change (e.g., [Von Oertzen, Hertzog, Linden](#page-21-17)[berger, & Ghisletta, 2010\)](#page-21-17). One way to think of the power to detect differences in cognitive change is to assume that there are two levels of the potential correlate, with equal numbers of participants at each level. Within a framework such as this it is possible to determine the number of participants in each group necessary to detect a given effect size for a difference in change with a specified degree of power. As an example, the sample size needed to achieve .8 power with a two-tailed significance level of .01 for a medium (i.e., Cohen's *d* of .5) effect on change is 96 per group, and 586 participants per group would be needed to detect a small (i.e., Cohen's *d* of .2) effect on change.

In addition to the size of the sample, characteristics of the participants in the sample are also important. For example, if individuals with cognitive impairments, either at baseline or emerging during the course of the longitudinal evaluation, are included in the analyses, the results may reflect effects of disease processes as much or more than effects of normal aging. These individuals are obviously interesting for other questions, but their inclusion could distort inferences about what occurs in healthy aging.

It is also desirable to obtain information about the representativeness of the initial sample as well as the selectivity of attrition in the longitudinal sample. Individuals who continue to participate in longitudinal studies frequently have higher scores at the initial occasion than individuals who do not continue, and this can affect the generalizability of the results [\(Salthouse, in press-b\)](#page-20-25). Selective attrition is also an important consideration in analyses of change because estimates of change, and correlates of change, can be distorted if selective attrition results in a restriction in the range of variation of the potential correlate or of the measures of cognitive functioning.

Finally, much of the prior research concerned with correlates of change has involved adults over about 65 years of age, and thus relatively little information is available about correlates of cognitive change at younger ages. This is unfortunate because different patterns might be expected at different ages if increased age is associated with shifts in the direction, magnitude, or causes of cognitive change. For example, different correlations of change might be expected at different ages if the change is primarily positive at young ages because of greater retest effects and is primarily negative at older ages because of greater maturationrelated effects. It is also possible that influences accumulate over time, such that the effects of the correlate are pronounced only at older ages.

Present Study

The current project incorporated the characteristics just described in an investigation of correlates of cognitive change. The initial sample consisted of a total of 4,802 adults, of whom 2,265 returned for a second measurement occasion, and 1,128 returned for a third measurement occasion. Longitudinal change was examined in five cognitive abilities, with each ability represented by either three or four different tests. Thirty potential correlates of change ranging from measures of sensory ability to aspects of lifestyle were examined both independently and in simultaneous analyses. Because it may not be meaningful to study change as a quantitative phenomenon if the nature of the construct shifts from one occasion to the next, longitudinal measurement invariance for each cognitive ability construct was examined first. Means and variances of the latent level and latent change parameters were next examined among adults between 65 and 99 years of age and among adults between 18 and 39 and between 40 and 64 years of age. The former group corresponds to the typical age range of prior studies in which correlates of change have been reported, and the latter two groups allow the comparisons to be extended to younger ages. Cognitive change was analyzed with second-order (sometimes referred to as multiple-indicator) latent growth curve models in which the latent level and latent change constructs correspond to the second level, with latent constructs based on three or four variables for each cognitive ability at each occasion representing the first level (cf. [Figure 1\)](#page-3-0). Finally, relations of potential correlates with latent level and latent change parameters were examined in the three age groups.

Method

Participants

Research participants were recruited from newspaper advertisements, flyers, and referrals from other participants. Approximately 81% of the participants were Caucasian, about 11% African American, and the remainder distributed across other ethnicities or reporting more than one ethnicity. Demographic characteristics of the participants in the three age groups as a function of number of occasions are summarized in [Table 1,](#page-4-0) with the right-most column containing the differences between numbers of occasions expressed in *d* units of effect size. The correlations of the demographic variables and composite cognitive ability scores with age are also reported in the table. It can be seen that increased age was associated with slightly poorer self-ratings of health but higher levels of education.

Figure 1. Illustration of the second-order latent growth model used in the analyses of longitudinal change. Unlabeled paths were freely estimated, and others were either constrained to the specified value or to be the same for relations with the same label. The paths in dotted lines represent the influences of the predictors on the latent level and latent change variables. Lvl = level; Chng = change; $T1-T3 =$ Time 1–Time 3.

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Note. Numbers in parentheses are standard deviations. Self-rated health was rated on a scale from 1 (*excellent*) to 5 (*poor*), and health-related limitations was rated on a scale from 1 (*none*) to 5 (*a great deal*). T1 to T3 = Time 1 to Time 3.
^a Indicates that the mean difference was significant at $p < .01$.

 $p < .01$.

Representativeness

In a recent study [\(Salthouse, in press-a\)](#page-20-26), both the present test battery and the Wechsler Adult Intelligence Scale IV [\(Wechsler,](#page-21-18) [2008\)](#page-21-18) test battery were completed by 90 adults between 20 and 80 years of age, which allowed estimates of full scale IQ scores to be derived in the current participants. Because IQ scores are ageadjusted, the estimation procedure consisted of partialing age from the raw scores to create residual scores, determining the best prediction of IQ from the residual scores, and then using the resulting regression equation to estimate IQ in the sample of 90 adults who performed both batteries. The most parsimonious regression equation with good prediction of IQ (i.e., $R^2 = .86$) was $109.32 + 2.47$ (series completion residual) $+ 1.54$ (antonym vocabulary residual) $+1.78$ (paper folding residual). This equation was applied to all of the current participants with relevant data to generate estimated IQ values.

Selective Attrition

The Virginia Cognitive Aging Project is an ongoing longitudinal study in which new participants are recruited each year and prior participants are invited to return after an average interval of about 3 years. Because approximately 800 of the individuals in the present sample participated for the first time within the last 3 years, they have not yet been invited to return for a second occasion. More information on the reasons for the attrition among the eligible participants is reported in [Salthouse \(in press-b\).](#page-20-25)

The data in [Table 1](#page-4-0) are informative about the selectivity of the longitudinal participants relevant to the initial sample. Among the adults between 18 and 39 years of age, participants with more occasions were older than participants with fewer occasions, but the reverse was the case for adults between 65 and 99 years of age. There were relatively small differences in self-rated health and years of education associated with number of occasions, but participants in the two older groups with two or more occasions had higher estimated IQs and composite cognitive ability scores at the first occasion than did participants with only one occasion. This pattern was reversed in participants between 18 and 39 years of age, which is likely attributable to greater mobility among the highest ability young individuals.

Although the participants who returned on subsequent occasions tended to have higher levels of cognitive performance on the first occasion than those who did not return, it is important to note that this does not necessarily limit the generalizability of the results regarding correlates of change. That is, selective attrition would not necessarily affect generalizability if the magnitude of longitu-

dinal change was similar across different levels of initial ability. In fact, little or no relations between initial ability and magnitude of longitudinal change were reported by [Salthouse \(2012\)](#page-20-27) after controlling influences associated with regression toward the mean, and [Salthouse \(in press-b\)](#page-20-25) recently found that the estimates of imputed change for participants who did not return for a second occasion were similar to the observed values for participants who did return. In addition, analyses conducted using the present data revealed no significant differences between participants with two or three occasions on the magnitude of change from the first to the second occasion. That is, between-groups *t* tests were conducted on the composite score differences from Time 1 (T1) to T2 in each cognitive domain, and all of the *t* test values comparing participants with two or three occasions were less than 1.3, with effect sizes (in *d* units) ranging from .00 to .03.

Change Analyses

Change was analyzed with the second-order latent growth model portrayed in [Figure 1.](#page-3-0) The boxes in the figure represent measured (manifest) variables, and the circles represent unmeasured (latent) variables. Some of the latent variables represent the level at each occasion (T1, T2, and T3), others represent the level (Lvl) or change (Chng) across occasions, and still others represent residual (unexplained) variance. The possibility of variablespecific change was accommodated by specifying covariances among the residuals at each occasion for a given variable. It should be noted that the Lvl construct is determined equally by performance in all three occasions, whereas the Chng construct was determined progressively more by scores on later occasions. The basis coefficients for the three occasions representing the latent change variable were set to 0 and 1 for the first and third occasions, respectively, with the coefficient for the middle occasion estimated from the data.

Advantages of the model in [Figure 1](#page-3-0) over other methods of analyzing change are that the latent variables representing level and change theoretically have no measurement error because only systematic variance can be shared, and estimates of means and variances of the level and change variables, as well as the relations between them, are available. Furthermore, rather than only analyzing data from individuals with complete data, missing data were handled by assuming that the data were missing at random and using the full information maximum likelihood (FIML) algorithm in the AMOS [\(Arbuckle, 2007\)](#page-19-19) modeling program. The FIML procedure uses all available data in the analyses, which not only increases precision and yields less biased estimates than analyses based on complete cases, but by including data from individuals only tested once, it also provides some adjustment for longitudinal selectivity. Unlike imputation procedures, in which estimates of the missing data are first derived and then the analyses conducted on the combined original and imputed data, FIML procedures handle the missing data and estimate parameters and standard errors in a single step [\(Graham, 2009;](#page-19-20) [Schafer & Graham, 2002\)](#page-21-19).

Because estimates based on participants with complete data could differ from those based on the FIML procedure, the same types of latent change analyses were also conducted on the sample of 1,128 participants who had data on all three longitudinal occasions. Although these analyses were less powerful than the primary analyses because of the smaller sample size, the results were very similar to those in the primary analyses. For example, of the 306 possible predictors of cognitive change (nine each for the 30 variables and four factors), 11 were significant in the FIML analyses in [Tables 7](#page-13-0) and [9,](#page-17-0) and nine were significant in the parallel analyses of the sample with data on all three occasions.

Cognitive Variables

Cognitive functioning was assessed with 16 tests selected to represent five cognitive abilities: word knowledge (vocabulary), inductive reasoning (reasoning), spatial visualization (space), episodic memory (memory), and perceptual speed (speed). Identical test versions were used at each longitudinal occasion. All of the individual test variables had coefficient alpha and test–retest reliabilities of .7 or higher and loadings of .7 or greater on their respective ability factors. The measures are briefly described in the [Appendix,](#page-22-0) and more details, including sources of the tests, are contained in other publications (e.g., [Salthouse, 2004;](#page-20-28) [Salthouse &](#page-20-29) [Ferrer-Caja, 2003;](#page-20-29) [Salthouse, Pink, & Tucker-Drob, 2008\)](#page-21-20). Scores at each occasion were converted to *z* scores based on the means and standard deviations of the scores at the first occasion. Ability constructs were formed at each occasion from the three or four (for vocabulary) measures established to have high loadings on the relevant ability factor. For some analyses, composite scores were created by averaging the *z* scores for the measures representing each ability.

Potential Correlates of Change

Self-rated health was assessed with two questions: "How would you rate your health at the current time?" on a scale from 1 (*Excellent*) to 5 (*Poor*) and "How much are your daily activities limited in any way by your health or health-related problems?" on a scale from 1 (*Not at all*) to 5 (*A great deal*). Visual acuity in both the right and left eyes was assessed with the Lighthouse Near Visual Acuity Test while the participants were wearing any prescribed corrective lenses. The denominator of the Snellen ratio was used as the measure of acuity.

Additional questions asked participants to evaluate their own memory and thinking abilities. The memory rating was the average of three ratings of memory compared to the average individual, to the best it has ever been, and in terms of problems experienced, on scales from 1 (*Very poor* or *Much worse*) to 7 (*Very good* or *Much better*). The thinking rating was the average of two ratings of thinking and reasoning relative to earlier in life and in terms of problems in day-to-day life on scales from 1 (*Much worse* or *Interferes a lot*) to 7 (*Much better* or *Does not interfere*).

Other potential correlates were obtained from questionnaires completed by the participants at home. Depressive symptoms were assessed with the Center for Epidemiological Studies–Depression scale [\(Radloff, 1977\)](#page-20-30), and trait anxiety was assessed with the Spielberger State–Trait Anxiety Inventory [\(Spielberger, Gorsuch,](#page-21-21) [Lushene, Vagg, & Jacobs, 1983\)](#page-21-21). The Big Five personality traits were assessed with the International Personality Item Pool questionnaire [\(Goldberg, 1999;](#page-19-21) 50-item version). Dispositions were assessed with the Satisfaction With Life Scale [\(Diener, Emmons,](#page-19-22) [Larsen, & Griffin, 1985\)](#page-19-22) and with the 18-item version of the Need for Cognition Questionnaire [\(Cacioppo, Petty, Feinstein, & Jarvis,](#page-19-23) [1996\)](#page-19-23). Mood was assessed with the Positive and Negative Affect Scale [\(Watson, Clark, & Tellegen, 1988\)](#page-21-22), and self-reported problems with executive functioning were assessed with the Dysexecutive Questionnaire [\(Wilson, Alderman, Burgess, Emslie, &](#page-21-23) [Evans, 1996\)](#page-21-23). The [Martin and Park \(2003\)](#page-20-31) Busyness scale was completed to assess self-perceived busyness and routineness of one's lifestyle.

Two locally developed questionnaires were designed to assess aspects of lifestyle related to cognitive and physical activity. The cognitive activity questionnaire [\(Salthouse et al., 2002\)](#page-20-22) asked participants to indicate the number of hours they devoted to each of 22 activities and to rate the cognitive demands of the activities. The measure of cognitive activity used in the analyses was the number of hours per week devoted to the seven activities with the highest average ratings of cognitive demands (i.e., reading newspapers, using a computer, driving a car, reading nonfiction, working crossword puzzles, handling finances, and writing).

In an attempt to increase the validity of the self-reports of physical activity, items in the physical activity questionnaire asked about the number of times per month and the duration each time engaged in specific activities (i.e., walking, yard work, calisthenics, running, aerobics, swimming, tennis, rowing, cycling, and sports). The participants were also given an opportunity to list other activities, and among those mentioned were weight lifting, yoga, dance, and sex. However, because they were not systematically assessed from everyone, these other activities were not included in the present analyses. The measure of physical activity for each primary activity was the estimated hours per month, derived by multiplying the frequency per month by the time at each occasion. In addition, the total number of hours per month engaged in all activities, derived by summing the hours in the 10 specified activities, was used as an additional measure of physical activity.¹

Results

Composite Scores Across Occasions

Composite scores at each occasion for participants with complete data for different numbers of occasions are plotted in [Figure](#page-7-0) [2](#page-7-0) for memory and in [Figure 3](#page-7-1) for the other cognitive domains. Notice that the values were lower with increased age for each cognitive domain except vocabulary. Consistent with the selective attrition results, with the exception of the youngest group, the means were higher for participants with more occasions. It can also

¹ Because of the relatively large sample sizes, a significance level of .01 was used to determine statistical significance.

Figure 2. Means (and standard errors) of the composite memory score at each occasion for participants with one, two, or three occasions in adults in three age groups. $T1 =$ Time 1; $SD =$ standard deviation.

be seen that the lines connecting the means across successive occasions were flat for reasoning and space in the older group, which suggests that there was little mean change in these cognitive abilities for adults in the sample over 65 years of age.

Reliability and Stability

An initial set of analyses examined properties of the cognitive variables at different levels of aggregation to determine the level that might be most meaningful in the analyses of change. For each individual variable, composite variable, and latent variable, correlations were computed between scores on two sessions in the first occasion as an estimate of immediate test–retest reliability, and between the first and third occasion as an estimate of long-term (approximately six years) stability. Data reported in [Salthouse and](#page-21-24) [Tucker-Drob \(2008\)](#page-21-24) were used to compute the short-term retest correlations because in that study 56 participants between 18 and 39 years of age, 113 participants between 40 and 64 years of age, and 58 participants between 65 and 99 years of age performed identical versions of the tests on a second session approximately one week after the initial session.

Correlations from these analyses are reported in [Table 2,](#page-8-0) where it can be seen that most of the reliabilities were above .70 and were similar in the three age groups. The reliabilities were higher for composite scores than for individual variables and were highest for latent variables. Estimates of stability from the first to the third occasion were lowest for individual variables and highest for latent variables, but unlike reliabilities, the stabilities were generally lower for participants ages 65 to 99 than for participants in the younger groups.

Figure 3. Means (and standard errors) of the composite scores in four cognitive domains at each occasion for participants with one, two, or three occasions in three age groups. $T1 =$ Time 1; $SD =$ standard deviation.

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Table 2

Variable Ages 18–39 Ages 40–64 Ages 65–99 Rel *r*13 D13Rel Rel *r*13 D13Rel Rel *r*13 D13Rel Memory Word recall .64 .66 -0.06 .78 .55 .51 .51 .51 .57 .51 Paired associates **10** .71 .71 .00 .73 .61 .31 .78 .61 .44
 Logical memory .79 .68 .34 .79 .61 .46 .72 .54 .39 Logical memory .79 .68 .34 .79 .61 .46 .72 .54 .39 Composite .84 .81 .16 .86 .71 .52 .86 .68 .56 81. 1.1ent variable .1. .99 .89 .91 .97 .85 .80 .99 .89 .89 .89 .89 .85 Speed Digit symbol .85 .81 .21 .88 .74 .54 .87 .68 .59 Pattern comparison and a sample of the set of t
11 .52 .40 .50 .50 .59 .63 .59 .46 .78 .59 .46 .59 .63 .63 Letter comparison .80 .63 .46 .78 .59 .46 .85 .59 .63
Composite .92 .77 .65 .88 .74 .54 .91 .69 .71 71. 169 .09 .02 .77 .65 .88 .74 .54 .91 .69 .77 1.00 .89 1.00 .85 .73 .95 .81 .74 Vocabulary WAIS-IV vocabulary .90 .84 .38 .91 .74 .65 .85 .57 .65 74. 90. 91. 1.86 .36 .92 1.86 1.91 .96 .96 .96 .96 .99 .99 .99 .99 .99 .99 .99 Synonym .86 .78 .36 .84 .85 .07 .85 .69 .52 23. 70. 77. 10. 79. 10. 81 .51 .61 .61 .61 .52 .52 Antonym 79 .76 .92 .95 .95 .95 .94 .92 .25 .95 .79 .76 Latent variable .99 .98 .50 .99 .98 .50 1.00 .90 1.00 Reasoning Matrix reasoning .83 .75 .32 .75 .66 .26 .62 .64 -.06

Shipley abstraction .83 .82 .06 .87 .80 .35 .78 .70 .27 Shipley abstraction .83 .82 .06 .87 .80 .35 .78 .70 .27

Letter sets .51 .69 .65 .11 .17 .83 .65 .51 .69 .65 .11 1. 17 .17 .69 .65 .11 .17 .83 .65 .11 .17 .17 .17 .17 .17 .15 .69 .65 11. x6 .84 .82 .89 .87 .15 .15 .91 .86 .36 .84 .82 .11 Latent variable .96 .93 .43 .97 .95 .40 .97 .94 .50 Space Spatial relations .74 .88 − 1.17 .82 .81 .05 .81 .71 .34 Paper folding .79 .78 .05 .72 .65 .20 .65 .54 .24 Form boards .84 .76 .33 .77 .62 .39 .59 .59 .00 28. 175 .28 .28 .18 .89 .18 .89 .84 .31 .82 .75 .28 Latent variable 1.00 .95 1.00 1.00 .95 1.00 .95 .87 .62 Medians 19 Individual variables 28 .83 .76 .33 .80 .66 .38 .79 .63 .41
18 .89 .84 .36 .86 .75 .56 .56 .56 .56 Composite variables Latent variables .99 .93 .91 .97 .95 .73 .96 .87 .74

Estimates of First Occasion Reliability (Rel), Stability Correlation Between the First and Third Occasions (r13), and Reliability of the Difference Between the First and Third Occasions (D13Rel) for Individual, Composite, and Latent Variables in Three Age Groups

Note. Rel refers to test–retest reliability over a period of about 1 week with data from [Salthouse and Tucker-Drob \(2008\),](#page-21-24) *r*13 is the (stability) correlation between scores on the first and third longitudinal occasions, and D13Rel is the estimated reliability of the difference between the Time 1 and Time 3 scores as computed from $(Rel - r13)/(1 - r13)$. WAIS-IV = Wechsler Adult Intelligence Scale IV [\(Wechsler, 2008\)](#page-21-18).

Stability is inversely related to amount of change, and therefore high stability implies small individual differences in change. However, low stability does not necessarily imply large individual differences in change, because reliability also needs to be considered when interpreting the stabilities. One method of incorporating both reliability and stability information involves estimating the reliability of the difference between scores on the first and third occasion with the formula

 $r_{\text{difference from 1 to 3}} = (\text{reliability} - \text{stability}_{1-3})$ $\sqrt{(1 - \text{stability}_{1-3})},$

assuming equal reliabilities at each occasion (see [Cohen & Cohen,](#page-19-24) [1983,](#page-19-24) p. 69). Because the stability coefficients indicate the proportion of variance in the T3 score shared with the T1 score, one minus the stability coefficient indicates the proportion of T3 variance not predicted from the initial score that could be associated with change. The formula can therefore be interpreted as providing

an estimate of the proportion of reliable variance at T3 that is potentially attributable to change.

Although the values can be considered only approximations, the estimates of difference score reliability are clearly much higher for latent variables than for either individual variables or composite variables. Even though composite variables are aggregates and have higher reliability than individual variables, the estimated reliabilities of the T1 to T3 differences for composite scores were modest. Because they had the highest reliabilities at the initial occasion, as well as for the T1 to T3 differences, the subsequent analyses focused on latent variables.

Measurement Invariance Across Longitudinal Occasions

Longitudinal measurement invariance (e.g., [Ferrer, Balluerka, &](#page-19-25) [Widaman, 2008\)](#page-19-25) across the three occasions was examined separately for each cognitive domain in the three age groups. The analyses were based on latent variables with either three or four (for vocabulary) manifest variables at each occasion. Model 1 corresponds to configural invariance in which there were acrosstime correlations of the factors and of the residuals for each variable but no constraints on the parameter estimates at each occasion. Model 2 corresponds to weak factor invariance and differs from Model 1 in that the factor loadings were constrained to be equal at each occasion. Model 3 corresponds to strong factor invariance and differs from Model 2 in that intercepts (means of the manifest variables) were also constrained to be equal across occasions. Finally, Model 4 corresponds to strict factor invariance and differs from Model 3 in that unique variances for the variables were also constrained to be equal at each occasion.

Results of the invariance analyses for each cognitive ability in the three age groups are reported in [Table 3.](#page-9-0) Values for Model 1 (configural invariance) are presented in the first row within each set, where it can be seen that this model had excellent fits to the data in each cognitive domain. The difference in the chi-square test indicated significant loss of fit when progressively more constraints were imposed, particularly when intercepts of the observed variables were specified to be equal across time (Model 3). However, it is important to note that the absolute fit was quite good (i.e., comparative fit index $[CFI] > .95$, root-mean-square error of approximation $[RMSEA] < .05$ for all models, including the strict factor invariance model incorporating all constraints. It therefore seems reasonable to conclude that although the measurement properties of the cognitive ability constructs were not identical across occasions, they were nevertheless very similar.

Measurement of Level and Change

The latent growth model portrayed in [Figure 1](#page-3-0) was fit to the data with each cognitive ability in each age group. Fit statistics (reported in the first three columns of [Table 4\)](#page-10-0) with all combinations of abilities and age groups indicated that the model had excellent fits to the data, with all CFIs $> .98$ and RMSEAs $< .06$, and medians of .99 and .02, respectively.

[Table 4](#page-10-0) also contains estimated means and variances of the latent level and latent change variables for the five abilities in the three age groups. The estimated standard errors were converted to standard deviations to allow computation of *d* values of effect sizes for the age group differences. As expected, there were large age differences in the level estimates, with progressively lower means at older ages for all cognitive abilities except vocabulary, where the direction of the age difference was reversed. The variances of the level estimates were similar across age groups, with the exception of larger values for reasoning and smaller values for space at older ages. All of the change estimates were more negative at older ages, with significant positive change in every ability in the 18–39 group and significant negative change in memory, speed, and vocabulary in the 65–99 group.

The variances of the change estimates were small compared to the variances of the level estimates, but in the older group only the values for reasoning and space abilities were not significantly greater than zero. The estimates of change variance in memory and vocabulary were significantly larger in the 65–99 group than in the younger groups.

Table 3

Note. CFI values greater than about .90 and RMSEA values less than about .08 are often considered to reflect a reasonably good fit [\(Kline, 2005\)](#page-20-32). Model 1 is configural invariance, Model 2 is weak factor invariance (equal factor loadings), Model 3 is strong factor invariance (equal intercepts), and Model 4 is strict factor invariance (equal unique variances). CFI = comparative fit index; RMSEA = root-mean-square error of approximation.

Note. Estimates of effect sizes (d) of the group difference are derived from the standard errors of the unstandardized coefficients. CFI = comparative fit index; RMSEA = root-mean-square error of approximation; Var. = variation; T2 = Time 2; Y = young; M = middle-aged; O = old.
^a Indicates that the difference in raw regression coefficients was significant at $p < .01$. $p < .01$.

Entries in the column labeled Change T2 are estimates of the basis coefficients representing the proportion of the interval between T1 and T3 that provided the best fit for a growth function. Most of the values were between about .2 and .7, indicating nearly equal change in the two intervals (T1 to T2 and T2 to T3). However, the coefficients for reasoning and space in the older group were small or negative rather than positive, which suggests that change in these domains may not have been systematic for participants 65 years and older.

Finally, the last column contains level–change relations. Nearly all of the estimates were small, and thus there was little evidence in these analyses that the magnitude of change was related to the level of that ability.

It is noteworthy that there was no significant change variance in the reasoning and space domains in any of the three age groups. These results are consistent with the very high stabilities and low estimated reliabilities of the 1–3 differences in [Table 2.](#page-8-0) Because correlates of change cannot be expected when there is little systematic variance in change, the reasoning and space ability measures were not included in subsequent analyses.

Power Analyses

Statistical power was computed with the method outlined in the introduction in which the possible correlates were considered to be dichotomous and a two-group contrast was specified with sample sizes equal to one half of the sample in each group (i.e., $N = 656$ in the $18-39$ group, $N = 1,165$ in the $40-64$ group, and $N = 579$ in the 65–99 group). The analyses revealed that the power to detect a medium $(d = 0.5)$ effect size with a two-tailed test and a significance level of .01 was 1.0 in each group, and the power to detect a small $(d = 0.2)$ effect size was $.85$ in the 18–39 group, .99 in the 40–64 group, and .79 in the 65–99 group. In order to place this information in context, estimates of effect sizes were computed for a difference corresponding to 50% of the observed mean change. These effect sizes, and the corresponding power to detect a difference of that magnitude as significant (two-tailed alpha of .01), were .90 and 1.00 for memory in the 18–39 group, .15 and .85 for memory in the $40-64$ group, $-.19$ and .74 for memory in the 65–99 group, .16 and .62 for speed in the $18-39$ group, $-.12$

Table 5 *Summary Statistics for Potential Moderators of Cognitive Change*

and .62 for speed in the $40-64$ group, $-.22$ and .88 for speed in the 65–99 group, .61 and 1.00 for vocabulary in the 18–39 group, .14 and .79 for vocabulary in the $40-64$ group, and -15 and .49 for vocabulary in the 65–99 group. Note that because the mean changes were small and were associated with moderate variability, even a substantial difference equal to one half of the observed change corresponds to a small effect size. Nevertheless, even with these small effect sizes, the statistical power in the present study was greater than .74 for the memory changes in all three groups and for the change in speed in the older group, and it was greater than .60 for all except change in vocabulary in the oldest group.

Analyses of Potential Correlates

[Table 5](#page-11-0) contains means, standard deviations, coefficient alphas, stability coefficients between the first and third occasion, and linear and quadratic age relations for each potential correlate. All coefficient alphas except that for self-rated health were above .7, indicating good internal consistency. No internal consistency values are reported for the cognitive activity measures because the total score is based on different types of activities (e.g., using a computer and driving a car),

Variable *N* at T1 *M* SD α T1–T3 Corr.

are in hours per month. Numbers in parentheses for physical activity variables are the percentages of participants with nonzero values for that activity. T1–T3 = Time 1 to Time 3; Corr. = correlation; Std. coeff. = standard coefficient; CES-D = Center for Epidemiologic Studies Depression Scale; PANAS-Positive = Positive subscale of the Positive and Negative Affect Schedule; PANAS-Negative = Negative subscale of the PANAS.
* $n \leq 0$ $p < .01$.

Std. coeff. Age Age^2 which could be inversely related to one another, or for the physical activity measures based on single scores.

The estimates of stability from T1 to T3 were modest for self-rated health and objectively assessed visual acuity, PANAS negative mood, and self-rated thinking, but were above .55 for most other variables. The stability coefficients for the measures of cognitive and physical activities were very low, indicating little consistency from the first to the third occasion.

The age relations were generally as expected in that increased age was associated with poorer self-rated health and lower visual acuity (both indicated by higher numbers); lower self-ratings of depressive symptoms, anxiety, and negative mood but higher self-ratings of positive mood and emotional stability; lower reported busyness; higher reported routine; and poorer ratings of one's level of memory and thinking. Some quadratic age trends were significant, indicating acceleration of the age relations if the quadratic trend had the same sign as the linear trend and flattening of the age relations if the sign was in the opposite direction.

Latent growth models were examined with each potential correlate to determine if there was significant variance in the change in the correlate. Many of the estimates of change variance were not significantly greater than zero, which implies very small individual differences in change in the correlate. Because most of the potential correlates had moderately high stability coefficients, the value at T1 was used as the predictor of level and change in cognitive abilities in all subsequent analyses. Although this precludes potentially informative analyses of the relations of correlate change with cognitive change, the measures at the first occasion were more reliable than the measures of change. In addition, assessment of the correlate at the first occasion minimizes ambiguity about reciprocal causation because subsequent cognitive change is unlikely to be the cause of the initial value of the correlate.

Correlates of Level and Change

An initial set of analyses examined demographic characteristics at T1 as simultaneous predictors of the latent level and latent change estimates in each cognitive domain, and standardized coefficients from these analyses are reported in [Table 6.](#page-12-0)

With the exception of vocabulary ability in the 40–64 group, all of the relations of age with the level estimates were significantly negative, indicating lower levels at older ages. All of the relations of age with the latent change estimates were negative but were significant only for some of the comparisons. However, the *d* values indicate that the differences in the unstandardized coefficients relating age to change were relatively small.

Females had higher average scores than males in memory and also slightly higher levels of speed in the 40–64 group, and they had slightly lower levels of vocabulary in the $18-39$ group. Longitudinal change in memory was less negative for females than for males in the 65–99 group but not in the 18–39 group. More education was associated with higher levels of performance in each ability domain, although the relations were weaker in the 65–99 group than in either of the younger groups. Importantly, the only relation of education to change was more positive change in the 18–39 group.

Because most prior studies considered potential correlates of cognitive change in separate analyses, each potential correlate was initially examined individually with only age, sex, and education as covariates. Standardized coefficients from these analyses are presented in [Table 7.](#page-13-0)

Most of the relations on the level coefficients were as expected, with higher levels of cognitive performance associated with better health, better vision, fewer depressive symptoms, less negative mood, lower self-reported executive (Dysexecutive Questionnaire) problems, higher openness, higher agreeableness, higher need for cognition, higher life satisfaction, more cognitive activities, and higher self-ratings of memory and thinking. However, the significant negative relations between positive mood and both memory and vocabulary were unexpected, as were the weak relations with emotional stability (the reverse of neuroticism). Only the relations with cognitive activity exhibited much of a difference across age groups, with more positive relations of reported cognitive activity

Table 6

Standardized Relations of Demographic Predictors at T1 With Level and Change in Cognitive Abilities, and Effect Sizes

Variable	Level			\boldsymbol{d}			Change			\boldsymbol{d}		
	Y $18 - 39$	M $40 - 64$	Ω $65 - 99$	$Y-M$	$M-O$	$Y - O$	Y $18 - 39$	M $40 - 64$	Ω $65 - 99$	$Y-M$	$M-O$	$Y - O$
Age												
Memory	$-.31*$	$-.17*$	$-.47*$.15 ^a	$-.31a$	$-.10^{\rm a}$	$-.38$	$-.13$	$-.18*$.04	$-.06$	$-.01$
Speed	$-.29*$	$-.33*$	$-.49*$	$-.03$	$-.18^{\rm a}$	$-.20^{\rm a}$	$-.50^*$	$-.15$	$-.19*$.10 ^a	$-.04$.07
Vocabulary	$-.11$ [*]	$.13*$	$-.11*$.30 ^a	$-.30a$.00	$-.54*$	$-.24*$	$-.10$.07	.09	.07
Sex^b												
Memory	$.10*$	$.22*$	$.26*$.12 ^a	.03	.15 ^a	$-.27$.15	$.17*$.10 ^a	.04	.13 ^a
Speed	.06	$.15*$.07	.09 ^a	$-.08$.02	.08	.13	.02	.01	$-.03$	$-.01$
Vocabulary	$-.13*$.00	.03	.13 ^a	.03	.17 ^a	$-.06$	$-.04$.11	.00	.05	.05
Education												
Memory	$.44*$	$.44*$	$.15*$	$-.06$	$-.26^{\rm a}$	$-.31a$	$.55*$.02	.08	$-.13^{\rm a}$.04	$-.08$
Speed	$.41*$	$.35*$	$.15*$	$-.07$	$-.19^a$	$-.25^{\rm a}$.05	.03	.00.	$-.01$	$-.01$	$-.02$
Vocabulary	$.58*$	$.57*$	$.45*$	$-.01$	$-.24^{\rm a}$	$-.23^{\rm a}$	$-.06$	$-.16$	$-.19$	$-.02$	$-.05$	$-.06$

Note. 18–39, 40–64, and 65–99 refer to participants' age category in years. Estimates of effect sizes (*d*) of the group difference are derived from the standard errors of the unstandardized coefficients. T1 = Time 1; Y = young; M = middle-aged; O = standard errors of the unstandardized coefficients. T1 = Time 1; Y = young; M = middle-aged; O = old.
^a indicates that the difference in raw regression coefficients was significant at $p < .01$. ^b Male = 0; female = 1.

 $p < .01$.

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Table 7 *Standardized Relations of Predictors of Latent Level and Latent Change in Different Cognitive Abilities, Controlling for Age, Sex,*

(*table continues*)

Table 7 (*continued*)

Table 7 (*continued*)

Note. 18–39, 40–64, and 65–99 refer to participants' age category in years. Estimates of effect sizes (*d*) of the group difference are derived from the standard errors of the unstandardized coefficients. Y = young; M = middle-aged; O = old; CES-D = Center for Epidemiologic Studies Depression Scale; PANAS-Positive = Positive subscale of the Positive and Negative Affect Schedule; PANAS-Negative = Negative subscale of the PANAS.
^a Indicates that the difference in raw regression coefficients was significant at $p < .01$ p^* *p* \lt .01.

on the level of cognitive performance for the older age group in all three cognitive domains.

Only nine predictors of change in [Table 7](#page-13-0) (out of 270, corresponding to a proportion of .03) were significant at $p < .01$, with five in the 18–39 group, two in the 40–64 group, and two in the 65–99 group. Furthermore, the effect sizes indicating differences between age groups in relations of the correlates were small, with *d* values ranging from $-.12$ to .11.

Because it is unlikely that the potential correlates were all independent of one another, the possibility of meaningful clusters of variables was investigated with exploratory factor analysis (principal axis factoring with promax rotation) on all 29 variables (excluding the sum of physical activities measure). Ten eigenvalues were greater than 1, but only four factors in which the same variables had high loadings on the factors in all three age groups were interpretable. The factor analysis results are summarized in [Table 8.](#page-16-0)

The first factor can be labeled Negative Affect because the highest positive loadings were with anxiety, depressive symptoms, negative mood from the PANAS, and the dysexecutive score. The second, third, and fourth factors can be labeled Openness, Self-Efficacy, and Busyness, respectively, because those variables had the strongest loadings in each factor.

The four factor scores were next used as simultaneous predictors of the latent level and latent change variables in the memory, speed, and vocabulary domains, with age, sex, and education as control variables. Results of these analyses are presented in [Table 9.](#page-17-0)

The only consistent relations across all three age groups were those of the Self-Efficacy factor on the level of memory and vocabulary. Two other relations on the level parameters were significant, a negative relation of Negative Affect on vocabulary in the 18–39 group and a negative relation of Busyness on vocabulary in the 40–64 group. Of primary interest were the

Openness Factor 2

Self-Efficacy Factor 3

* $.00$ $-.25$ $-.37^*$

 $.40^*$ $.33^*$
- $.64^*$

Busy Factor 4

 $-.37$ ^{*}

* $-.17^*$

* $-.06^*$

lary in the 65–99 group.

Table 8 *Exploratory Factor Analysis (Principal Axis Factoring, 4 Factors, Promax Rotation) on 29 Potential Correlates*

Negative Affect Factor 1

Discussion

positive relation between Self-Efficacy and change in vocabu-

As noted in the introduction, prior research on correlates of cognitive change has been inconsistent, and thus it is important that additional research on this topic be as methodologically rigorous as possible. The present study has a number of strengths compared to prior studies in which correlates of cognitive change have been investigated. For example, cognitive functioning was assessed at the level of latent variables defined by scores on three or four separate cognitive tests, which increases the breadth of assessment and minimizes measurement error relative to assessments with single variables. In addition, measurement invariance analyses indicated that the cognitive ability constructs had similar meaning at each measurement occasion, which implies that the changes were primarily quantitative rather than qualitative. Fur-

thermore, sensitive assessment of cognitive change based on a second-order latent growth model revealed that there was significant mean change, and significant variance in change, in the speed measure in all three age groups, in the memory measure in the 40–64 and 65–99 groups and in the vocabulary measure in the 65–99 group (cf. [Table 4\)](#page-10-0). Because there was no significant change variance in the measures of reasoning and space, and high stability from the first to the third occasion, it was not meaningful to examine correlates of change in those abilities in the present study. In addition, the statistical power to detect small differences in change was above .74 for the measures of memory in all three age groups, for the vocabulary measure in the 40–64 and 65–99 groups and for the speed measure in the older group. Most of the potential correlates had good coefficient alpha reliability, and in some analyses they were aggregated into factors that can be expected to be even more reliable. Although the sample of participants had a higher average level of functioning than that in a nationally representative normative sample, the magnitude of variability was similar, and there was little attenuation of the variabil-

Table 9 *Simultaneous Relations (Standardized) of Predictors of Latent Level and Latent Change in Different Cognitive Abilities, Controlling for Age, Sex, and Education*

Note. 18–39, 40–64, and 65–99 refer to participants' age category in years. Estimates of effect sizes (*d*) of the group difference are derived from the standard errors of the unstandardized coefficients. $Y =$ young; $M =$ middle-aged; $O =$ old; F1 to F4 = Factor 1 to Factor 4.
^a indicates that the difference in raw regression coefficients was significant at $p < .01$. p^* *p* \lt .01.

ity after attrition. In addition, unlike many earlier studies, in which all of the participants were over 65 years of age, the participants spanned a wide age range.

Many of the potential correlates had significant relations with measures of the level of cognitive functioning. As in other reports, there were negative relations of the measures of cognitive functioning with depressive symptoms, anxiety, and negative mood and positive relations with health, vision, openness, conscientiousness, need for cognition, life satisfaction, and self-ratings of memory and thinking. There were also a few age differences in the pattern of relations, as the female advantage for memory was greater in the two older groups and the effects of education were weaker in the oldest group.

The major results with both individual variables and with the factors representing groups of related variables were the weak to nonexistent correlations of cognitive change. Only three of the potential correlates were related to individual differences in change in memory: two in the 18–39 group, reflecting more negative memory change with higher values of negative mood and more positive memory change with higher levels of cycling activity, and one in the 65–99 group, in which individuals with more negative ratings of their memory had more negative change in memory. Three predictors had significant relations with change in speed, and all were in the 40–64 group. Individuals with less time in yard work, greater time in calisthenics, and greater time in running had more positive change in speed. The only significant predictor of change in vocabulary was in the 65–99 group, in which, surprisingly, people with a greater reported time in cognitive activities had more negative change in vocabulary.

The patterns of relations were generally similar in the three age groups, and therefore there was little evidence that relations with cognitive functioning were restricted to the period of late adulthood. The primary exception was the cognitive activity variable, which had more positive relations with the levels of memory, speed, and vocabulary in the two older age groups. These results indicate that among the middle-aged and older adults, people with more engagement in cognitively stimulating activities had higher levels of cognitive functioning than did people with less engagement. Although these results are interesting, it is important to recognize that the causal direction of this relation is ambiguous because level of ability could have contributed to participation in stimulating activities rather than participation contributing to different levels of ability. Furthermore, the relations of cognitive activity were primarily apparent with measures of the level of functioning, and the only relation with change in functioning was negative; thus, there was no support for the hypothesis that engagement in cognitively stimulating activities alters the rate of change in cognitive ability.

Despite numerous strengths, the present study failed to identify significant moderators of cognitive change. It is therefore important to consider factors that might be contributing to the inconsistencies in research concerned with correlates of cognitive change. Although a definitive answer is not yet available, at least seven possibilities that might account for different patterns of results in studies investigating correlates of change are worth considering.

First, it is conceivable that the published literature is somewhat distorted because negative findings might have been less likely to have been published than positive findings. In addition, some of the positive outcomes that were reported could have been attributable to chance because not all studies adjusted the significance level for the number of statistical comparisons.

Second, a variety of cognitive measures have been included in the prior studies, and some of the differences in results may reflect effects on different aspects of cognitive functioning. There has also been considerable variation in the outcome variables, as some studies have focused on incidence of pathological conditions such as dementia whereas others have been concerned with continuous change in cognitive functioning in healthy adults.

Third, many different measures of potential correlates have been examined, and even when they were described with the same label, they may not have represented the same construct. For example, in some studies a subset of items from the original scales was used, which may not have had the same reliability (because there were fewer items) or validity (because all facets of the construct may not have been represented) as the original scale. Activities have sometimes been assessed with a very small number of items, which might not have been very reliable or valid, particularly when evaluated with self-reports. The assessments might also have differed qualitatively and not quantitatively, as they have ranged from evaluation of presence or absence to measures combining time and intensity in multiple activities. There has also been considerable variation with respect to when the correlate was assessed, such as current, past, recent, or cumulative across one's lifetime and with respect to characteristics such as intensity and frequency.

Fourth, there have been many differences in the composition of the samples across studies, including the range of ages and ability levels and the magnitude and selectivity of attrition. Furthermore, in some studies the distribution of individuals with different numbers of measurement occasions was highly skewed, which implies that the change estimates were heavily influenced by a very small number of individuals from the initial sample. Some prior studies may also have included substantial proportions of individuals in early stages of dementia or terminal decline, which could have resulted in more negative mean change and/or greater variance in change relative to studies with only healthy adults.

Fifth, different analytical methods have been used to assess cognitive change, and some of the analyses of change may have been influenced by the mean level of performance or by the relation of the correlate to the baseline scores [\(Glymour, Weuve,](#page-19-26) [Berkman, Kawachi, & Robins, 2005\)](#page-19-26). Furthermore, measurement equivalence was seldom examined to evaluate comparability of the cognitive constructs across different occasions.

Sixth, many of the analyses may have had low power to detect potentially interesting differences in cognitive change. To illustrate, the findings in the present study that a 50% difference in the change in memory corresponded to an effect size of only $-.19$ in the 65–99 group suggest that even large differences in cognitive change may be difficult to detect without very large sample sizes.

Importantly, few studies have reported whether there was significant variance in change, which is necessary to have correlations with other variables. Unfortunately, little information is currently available about the magnitude of change variance in longitudinal studies. However, it is noteworthy that one major study in which variance of cognitive change was examined over a period of 10 years found significant change variance in only four of 20 comparisons (i.e., five age groups with four cognitive measures each), and none of those was significant after eliminating participants

who died or developed dementia during the interval [\(de Frias,](#page-19-27) [Lövdén, Lindenberger, & Nilsson, 2007\)](#page-19-27). If there is no evidence of differential change, it is unrealistic to expect to identify correlates of differences that do not exist.

Finally, the fact that there have been very few, if any, exact replications with the same measures and analytical procedures reinforces concerns raised about the role of "flexibility in designs, definitions, outcomes, and analytical modes" [\(Ioannidis, 2005,](#page-20-33) p. 0698) and "researcher degrees of freedom" [\(Simmons, Nelson, &](#page-21-25) [Simonsohn, 2011,](#page-21-25) p. 1359) in contributing to false positive results. As an example, although many studies have investigated the relation between physical activity and cognitive change, [Salthouse](#page-20-21) [\(2010,](#page-20-21) p. 144) noted that the available studies differed in many respects, including the measures of cognitive functioning, the analytical procedures, and the methods used to assess physical activity, such as self-rating at baseline, change in self-rating, or objectively assessed fitness. It may therefore be misleading to suggest that the studies are reporting the same result when they had so few features in common. Because longitudinal studies are expensive and time consuming, exact replications with longitudinal studies are rare. Nevertheless, three approximations to replications in longitudinal studies should be encouraged because they can be informative in examining the robustness of correlates of cognitive change: (1) comparing results across different subsamples within the same study, such as the three age groups in the present study; (2) comparing results across different cohorts recruited in different years (e.g., [Small et al., 2012\)](#page-21-15); and (3) using common models to analyze similar variables in different data sets (e.g., [Hofer & Piccinin, 2009;](#page-19-28) [Lindwall et al., 2012;](#page-20-24) [Mitchell et al.,](#page-20-3) [2012\)](#page-20-3).

It is not yet clear which, if any, of the preceding characteristics may have contributed to the different patterns of results regarding correlates of cognitive change in healthy adults. However, it is important to note that the present study had moderately large samples; assessment of multiple cognitive abilities at the level of latent variables, which minimizes measurement error; reliable assessment of potential correlates that were significantly related to many measures of the level of cognitive functioning; and powerful analytical methods that revealed significant variance in the change in memory, speed, and vocabulary. Although these features should have contributed to sensitive detection of correlates of change, there was little evidence in this study that aspects of lifestyle, mood, or personality moderate longitudinal change in cognitive functioning among healthy adults.

Much of the interest in correlates of cognitive change has been motivated by an interest in identifying possible targets for intervention. In a recent review of risk factors for cognitive decline, [Plassman et al. \(2010,](#page-20-1) p. 189) concluded that "the current literature does not provide adequate evidence to make recommendations for interventions." The results of the present study reinforce this conclusion because no consistent correlates of cognitive change could be identified. It is also important to recognize that even if significant correlations were found, drawing causal inferences from correlations should be done with great caution, and in particular, one should avoid implications that altering a correlate will necessarily alter the trajectory of cognitive change. As an example, a recent study reported that married individuals exhibited less memory decline than individuals who were not married [\(Mousavi-](#page-20-34)[Nasab et al., 2012\)](#page-20-34), but even if this finding were confirmed in other studies, marriage should not necessarily be advocated as an intervention to minimize cognitive decline.

In conclusion, the results of the present study suggest that in healthy adults, increased age is associated with more negative change in several major cognitive abilities and that there are significant individual differences in cognitive change, particularly among adults age 65 and older. However, there was little evidence of moderators of cognitive change across cognitive abilities, different age groups, or different analytical methods (e.g., with individual predictors or factors, and in both the FIML analyses and the analyses based on participants with data on all three occasions). Until a consistently replicated pattern based on methodologically strong studies has been established, therefore, the most reasonable conclusion at the current time may be that if these variables do moderate the rate of cognitive change, the effects are likely to be quite small.

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(*Appendix follows*)

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Appendix

Description of Cognitive Variables

Vocabulary

Wechsler Adult Intelligence Scale IV [\(Wechsler, 2008\)](#page-21-18) vocabulary. Provide definitions of words.

Picture vocabulary. Name the pictured object.

Antonym vocabulary. Select the best antonym of the target word.

Synonym vocabulary. Select the best synonym of the target word.

Reasoning

Matrix reasoning. Determine which pattern best completes the missing cell in a matrix.

Shipley abstraction. Determine the words or numbers that are the best continuation of a sequence.

Letter sets. Identify which of five groups of letters is different from the others.

Spatial Visualization

Spatial relations. Determine the correspondence between a 3–D figure and alternative 2–D figures.

Paper folding. Determine the pattern of holes that would result from a sequence of folds and a punch through the folded paper.

Form boards. Determine which combinations of shapes are needed to fill a larger shape.

Memory

Logical memory. Recall idea units across three stories.

Word recall. Recall words across four trials of the same word list.

Paired associates. Recall response terms when presented with a stimulus item.

Speed

Digit symbol. Use a code table to write the correct symbol below each digit.

Letter comparison. Same/different comparison of pairs of letter strings.

Pattern comparison. Same/different comparison of pairs of line patterns.

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